Exploring the Dynamics of Inflation, Interest Rates, and Bond Yields: A Comprehensive Regression Model Comparison

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Abstract

This research presents a comprehensive analysis of the impact of key macroeconomic indicators, including 10year Government Securities (10Y GSec) yields, 91-day Treasury Bill (TB) rates, interest rates, inflation, exchange rates, foreign reserves, gold prices, equity market indices (NSE Nifty), FDI, and Month-on-Month (MoM) basis returns, on interest rates, inflation, yields and vice-versa in India, using data from 2018 to 2023. The study employs multivariate correlation analysis to identify the relationships among these variables, revealing significant patterns such as the strong correlation between bond yields and interest rates, as well as the inverse relationship between equity market performance and bond yields. This research highlights the influence of monetary policy, external reserves, inflation, FDI, and equity markets in shaping bond yields and interest rates, with gold and equity markets acting as safe-haven assets during times of economic uncertainty. A key focus of this research is a detailed comparison of five regression models— Ordinary Least Squares (OLS), Heteroscedasticity-Corrected (HSC), Tobit, Logistic, and the Transformed First Difference (FD) OLS—used to analyze the relationships among these macroeconomic variables. The study demonstrates that the HSC and Logistic regressions provide the most robust and reliable insights, with the HSC model exhibiting the highest explanatory power in capturing the variance in bond yields, interest rates, and inflation. The analysis underscores the importance of selecting the right regression model to accurately capture the dynamics of

financial markets, as model performance varies significantly. This comprehensive study offers a nuanced understanding of the interplay between macroeconomic indicators and financial outcomes, emphasizing the critical role of regression models in enhancing the accuracy of economic forecasting. **Keywords:**Ordinary Least Squares (OLS), Heteroscedasticity-Corrected (HSC), Tobit, Logistic, Transformed

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I. Introduction

This study delves into the complex relationships between inflation, interest rates, and bond yields, employing a diverse array of regression models to examine these macroeconomic variables. The models used in this research include Ordinary Least Squares (OLS), Heteroscedasticity-Corrected (HSC) regression, Tobit, Logistic, and the Transformed First Difference (FD) model. These models were chosen for their ability to address various econometric challenges. The HSC model corrects for heteroscedasticity, ensuring more accurate estimates when the variance of errors is not constant. The Tobit model addresses censored data, while the Logistic model is suitable for binary outcome analysis. The FD model, a form of transformation applied to panel data, removes individual-specific effects by differencing the data, thus improving the estimation of dynamic relationships in the context of panel datasets. By incorporating these diverse models, the study aims to provide a robust understanding of how inflation, interest rates, and bond yields are interlinked in both short-term and long-term scenarios.

The global bond markets operate within a complex web of economic interdependencies, where macroeconomic fundamentals and external shocks collectively shape yield dynamics. Inflation remains a pivotal factor, often displaying a positive correlation with bond yields, while gold and foreign exchange reserves tend to exhibit inverse relationships, reflecting their roles as safe-haven assets (Koroleva & Maxim, 2022). In emerging markets, factors like fiscal risk, demographic changes, and monetary policy innovations significantly contribute to yield variations, highlighting the diverse nature of bond market drivers (Michelson & Stein, 2023). Liquidity and market volatility are key considerations in understanding bond market behavior, particularly in times of financial crises. The asymmetric effects of liquidity risks on bond market stability (Sethy & Tripathy, 2024).

Similarly, sovereign bond spreads in emerging economies like Sri Lanka reveal a strong dependence on local fiscal and monetary policies, demonstrating the influence of domestic fundamentals amid global uncertainties (Kariyawasam & Jayasinghe, 2022).

Economic shocks, such as the COVID-19 pandemic, have reshaped the financial landscape by introducing unprecedented volatility and altering traditional market relationships. The pandemic revealed the fragile interconnectedness of global financial systems, as evidenced by the sharp declines in bond yields, stock prices, and capital flows in emerging market economies (Beirne et al., 2020). In India, COVID-19-induced volatility highlighted the differentiated impacts across industries and asset classes, illustrating the pandemic's heterogeneity in financial disruptions (Dharani et al., 2023). Capital flows and foreign portfolio investments are also integral to bond market dynamics, particularly in emerging markets with open economies. Research on India has shown that exchange rate volatility, interest rate differentials, and domestic growth significantly influence foreign portfolio flows, emphasizing the importance of stable macroeconomic policies (Gupta & Ahmed, 2019). In Peru, yield curve analyses reveal that while short- and medium-term yields are driven by observable macroeconomic factors, long-term movements are dominated by unobservable variables, indicating the need for robust modeling techniques to capture yield behavior accurately (Olivares Rios et al., 2019).

The relationship between bond yields and other financial markets, such as equity and commodity market, reflects broader economic conditions and market expectations. For instance, the sensitivity of bond yields to equity premiums in the U.S. demonstrates the interconnectedness of risk factors across asset classes (Bhar & Malliaris, 2011). Similarly, exchange rate dynamics play a crucial role in shaping bond yields, with currency movements significantly impacting international capital pricing and asset valuations (Francová, 2018). In emerging markets, global uncertainty indices and local financial stress indicators provide valuable insights into bond market behavior. For example, bond yield shocks in countries like Mexico, Russia, and South Korea reveal asymmetrical effects on exchange market pressures, suggesting the importance of local conditions in mitigating external financial stress (Ozcelebi, Perez-Montiel, & Manera, 2024). These findings underline the necessity for region-specific strategies in managing bond market risks.

Finally, advancements in econometric modeling have enhanced the understanding of bond market behavior, offering tools to analyze the impact of macroeconomic, financial, and behavioural factors. From wavelet analyses examining COVID-19 impacts on sovereign yields to ARDL approaches exploring foreign portfolio flows, these methodologies provide nuanced insights into yield determinants (Yilanci & Pata, 2023). Such approaches are crucial in navigating the multifaceted challenges of bond market analysis, enabling policymakers and investors to make informed decisions.

This study's use of multiple regression models is intended to provide a comprehensive view of the interactions between inflation, interest rates, and bond yields under varying economic conditions. By applying OLS, HSC, Tobit, Logistic, and FD models, the research aims to offer various perspectives on these relationships, from correcting for heteroscedasticity to managing censored data and analysing temporal changes in panel datasets. Each model is designed to address a different aspect of the data, thus enhancing the robustness of the study's findings. Ultimately, this research contributes to a deeper understanding of how key macroeconomic variables interact in today's rapidly evolving economic environment, offering valuable insights into economic forecasting and financial market behaviour.

II. Literature Review

The interplay between inflation, interest rates, and bond yields remains a cornerstone of financial market research, shaping both theoretical advancements and practical policy applications. Recent literature highlights the complexity and interdependence of these variables across developed and emerging markets. This review synthesizes insights from leading studies, emphasizing the nuanced dynamics influenced by macroeconomic factors, market structures, and global events. Literature reviews on the intricate relationships among inflation, interest rates, and bond yields often explore a range of econometric models and methodologies to elucidate these dynamics. This body of research provides critical insights into how macroeconomic factors influence financial markets across developed and emerging economies. By examining the findings of various studies, it becomes evident that a comparative perspective on methodologies such as OLS, Logit, and other advanced models offers valuable insights for understanding these relationships during normal and crisis periods.

In developed markets, Jiang et al. (2023) analysed the impact of key macroeconomic announcements on corporate bond markets, finding that good news negatively affected investment-grade bond returns while positively influencing high-yield bonds. Their study underscores the rapid incorporation of information into bond prices and heightened volatility during announcement days. Similarly, Zhou (2021) employed linear and non-linear ARDL models to investigate how government debt and macroeconomic variables impact South Africa's long-term bond yields, revealing that short-term interest rates are the most significant determinant of yields. Emerging markets display unique vulnerabilities and dynamics. For instance, Kalu et al. (2020) demonstrated how U.S. monetary policy normalization adversely impacts African stock prices due to global financial integration. Meanwhile, Moussa and Delhoumi (2022) used the NARDL model to assess the influence of interest and exchange rate fluctuations on stock returns in MENA countries, highlighting asymmetry in short-term effects. Nguyen and Nguyen (2022) identified inflation and policy rate changes as significant factors influencing Vietnam's government bond yields, with variations between short- and long-term bonds.

India offers a distinct perspective with its evolving financial markets. Tripathi and Seth (2014) explored the causal relationship between macroeconomic variables and stock market performance, identifying significant influences from inflation, interest rates, and exchange rates. Panigrahi et al. (2020) delved into the impact of macroeconomic variables on Indian mutual fund performance, finding that these factors account for over half of performance variability. Meanwhile, Garg and Kalra (2018) highlighted a positive correlation between the Sensex and macroeconomic indicators, except for inflation and unemployment, which negatively impacted the market. The COVID-19 pandemic introduced unprecedented shocks to financial systems worldwide. Hui and Chan (2022) revealed that global equity markets were significantly impacted, with European economies experiencing more pronounced effects than East Asian ones. Lakdawala et al. (2023) found that unconventional monetary policies by the Reserve Bank of India, such as liquidity support and asset purchases, effectively reduced bond yields, showcasing the critical role of central banks in stabilizing markets during crises. Zhou et al. (2022) compared Germany and the United States, finding divergent short-term responses to the pandemic's effects on government bond yields.

A comparison of econometric models highlights their utility in capturing different aspects of these complex relationships. Agrawal (2020) employed a Logit model to predict U.S. recession probabilities during COVID-19, focusing on indicators such as payroll and yield curve spreads. Flannery and Protopapadakis (2002) utilized GARCH models to examine stock market responses to macroeconomic variables, identifying CPI, PPI, and monetary aggregates as significant factors. Similarly, Çepni et al. (2022) used the Nelson-Siegel model combined with structural VAR to assess oil price shocks' impact on Turkey's yield curve. Studies on emerging economies during crises provide nuanced insights. Rabbani et al. (2024) used wavelet-based approaches to assess the influence of geopolitical risks on Islamic and composite stocks during the COVID-19 pandemic and the Russia–Ukraine conflict, finding significant impacts on market dynamics. Lebdaoui et al. (2024) employed E-GARCH analysis to study MENA markets, noting that economic resilience mitigated stock volatility during COVID-19. Sreenu and Pradhan (2023) examined sectoral volatility in India, identifying economic features that help stabilize markets under crisis conditions. Marisetty (2024) highlights the OLS Log Difference (OLS LD) model as the most effective in addressing regression assumption violations, including heteroscedasticity and autocorrelation.

Research in developed markets has also focused on structural changes. Takahashi (2022) used PCA to analyze yield curve movements across four major currencies, finding minimal structural changes pre- and post-2008 financial crisis. Chiang (2023) challenged the Fisher hypothesis by demonstrating a negative relationship between U.S. inflation and aggregate stock returns, with the energy sector as an exception. Comparative studies further enrich the discourse. Verma and Bansal (2021) analysed developed and emerging markets, finding that GDP, FDI, and FII positively influence stock markets, whereas gold prices and interest rates exert negative effects. Hughen and Beyer (2015) examined U.S. stock market performance, revealing that accommodative monetary policy and dollar appreciation drive returns, while depreciation and tight policies depress them. In smaller economies, macroeconomic variables often play outsized roles. Kyereboah-Coleman and Agyire-Tettey (2008) found that Ghana's lending rates and inflation negatively impacted stock market performance, while currency depreciation benefited investors. This highlights the vulnerabilities of smaller economies to macroeconomic fluctuations.

Sector-specific analyses also offer granular insights. Amin and Mollick (2022) explored the impact of oil prices on U.S. mining stocks, demonstrating that high leverage dampens the positive effects of price increases. Laila et al. (2021) examined corporate bond and sukuk ratings in Indonesia, identifying leverage and liquidity as critical determinants. On a global scale, interconnectedness among asset classes has been a recurring them. Rao et al. (2022) highlighted strong linkages between traditional and emerging asset classes, finding that Bitcoin and gold lost their safe-haven status under extreme economic conditions. Schrank (2024) focused on Thailand's monetary policy, revealing its significant influence on stock and bond markets during crises. Theoretical reviews complement empirical findings. Tangjitprom (2012) synthesized studies on macroeconomic factors and stock returns, categorizing variables into economic conditions, monetary policy, price levels, and international activities, emphasizing their varied but significant relationships.

In conclusion, the literature on inflation, interest rates, and bond yields highlights their intricate interrelationships across diverse economic contexts. Developed markets demonstrate structural stability, while emerging markets grapple with greater volatility and external dependencies. India's unique economic landscape provides valuable insights into these dynamics, particularly during global crises like COVID-19. Advanced modeling techniques and comprehensive sectoral analyses continue to enhance our understanding, offering actionable insights for policymakers and investors in an increasingly interconnected global financial system. This literature review examines how economic factors such as inflation, interest rates, exchange rates, and

equity markets affect bond yields and interest rates in India. It highlights the role of various regression models like OLS, HSC, Tobit, FD, and Logistic in studying these relationships. Recent studies emphasize the growing importance of inflation expectations and fiscal adjustments in determining long-term bond yields. Overall, the review shows how macroeconomic factors influence India's bond market, inflation, interest rates and how research methods continue to evolve.

III. Methodology

This study investigates the dynamic relationships among key macroeconomic variables in India using Month-on-Month (MoM) data from January 2018 to December 2023 and data collected from Reserve Bank of India (RBI). The dataset comprises 72 observations for each variable, including 10-year Government Securities (10Y GSec) yields, 91-day Treasury Bill (TB) rates, interest rates, inflation, exchange rates, foreign reserves, gold prices, NSE Nifty, and MoM returns. The analysis focuses on four dependent variables: 10Y GSec yields representing long-term market trends, 91-day TB rates capturing short-term monetary dynamics, interest rates as a reflection of borrowing costs, and inflation as a core economic indicator. These variables were selected for their central role in influencing financial stability and economic policy outcomes, serving as proxies for broader macroeconomic conditions.

The explanatory variables were chosen for their relevance to the dynamics of the dependent variables. Exchange rates, foreign reserves, and gold prices are key indicators of external sector strength and currency stability, which can significantly impact long-term yields and interest rates. NSE Nifty serves as a barometer for equity market performance, often linked to investor confidence and capital flows, while MoM returns capture short-term market volatilities. These diverse variables collectively provide a comprehensive framework for understanding the interplay between macroeconomic factors and financial markets, ensuring robust and multifaceted analysis.

A combination of regression models was employed to address the study's objectives, including Ordinary Least Squares (OLS), Heteroscedasticity-Corrected (HSC) regression, Tobit regression with left-censored values at 3% and right-censored values at 7%, Logistic regression for real-valued dependent variables such as interest rates and inflation, and Transformed First Difference (FD) OLS. These models were applied to accommodate data features such as censoring, non-linearity, and potential structural breaks, ensuring a nuanced exploration of the variables' relationships.

Rigorous diagnostic tests complemented the regression analyses to validate the models' reliability and stability. These included standard error evaluation, Log-Likelihood, Adjusted R-squared, and Akaike Information Criterion (AIC). Additionally, structural stability was assessed using the Chow test and QLR test for structural breaks, while the RESET test was used to verify the functional form of the models. Although the study provides a comprehensive framework for analysing macroeconomic dynamics, its limitations—such as the relatively short timeframe and inherent assumptions of some regression methods—underscore the importance of further research with extended datasets and alternative methodologies.

Descriptive statistics

Descriptive statistics were calculated for the nine variables, including bond yields, equity returns, and economic indicators, to summarize their central tendencies and variability. The study also used the Jarque-Bera test to assess the normality of the data distributions for each variable. The results of the test indicated whether the data significantly deviated from a normal distribution, guiding further statistical analysis. These preliminary steps provided a solid basis for the subsequent paired t-test and regression analyses.

Multivariate correlation analysis

Multivariate correlation analysis explores the relationships among multiple variables simultaneously. The correlation matrix is calculated using the formula:

$$\rho_{XY} = \frac{COV (X, Y)}{\sigma_X \sigma_Y}$$

where ρ_{XY} is the Pearson correlation coefficient, Cov (X,Y) is the covariance between variables X and Y, and σ_X and σ_Y are the standard deviations of X and Y, respectively. This analysis helps identify the strength and direction of relationships between bond yields, equity returns, and various economic factors.

Ordinary Least Squares (OLS) Regression

The Ordinary Least Squares (OLS) test in multiple regression estimates relationships between one dependent variable and multiple independent variables. The formula is:

$$Y = \beta_0 + \beta_1 X_1 + \beta_2 X_2 + \dots + \beta_n X_n + \epsilon$$

Here, Y is the dependent variable, X_1 , X_2, X_n are independent variables, β_0 is the intercept, β_1 β_{2} ..., β_{n} are coefficients, and ϵ is the error term. OLS minimizes the sum of squared residuals (ϵ^{2}) to estimate β values. Assumptions like linearity, no multicollinearity, and homoscedasticity are crucial for valid results.

Heteroscedasticity-Corrected (HSC) Regression

A Heteroscedasticity-CorrectedModel adjusts regression analyses to account for non-constant variance (heteroscedasticity) in the error terms, ensuring reliable estimates and valid statistical inference. The model corrects standard errors, often using robust techniques such as White's correction. The corrected regression equation remains:

 $Y = \beta_0 + \beta_1 X_1 + \beta_2 X_2 + \dots + \beta_n X_n + \varepsilon$

However, heteroscedasticity-adjusted standard errors are computed as:

 $\widehat{V} = (X'X)^{-1} X' \widehat{\Omega} X (X'X)^{-1}$

where is $\hat{\Omega}$ a diagonal matrix of error variances. This approach ensures unbiased coefficient estimates and accurate confidence intervals in the presence of heteroscedasticity.

Tobit Regression

Tobit regression is designed for datasets where the dependent variable (Y) is censored, either at a lower threshold (L) or an upper threshold (U). The model is expressed as:

$$Y_i^* = \beta_0 + \beta_1 X_{1i} + \beta_2 X_{2i} + \dots + \beta_k X_{ki} + \varepsilon_i, \varepsilon_i \sim N(0, \sigma^2)$$

Where: Y_i^* is the dependent variable. Y_i is the observed variable:

 $Y_i = \{L_i f Y_i \leq L, Y_i^* i f L < Y_i^* < U, U_i f Y_i^* \geq U \}$

- X_{ki} are the independent variables.
- β_i are the coefficients to be estimated, ϵ_i is the error term, normally distributed.

Logistic Regression

Logistic regression is used when the dependent variable is binary or categorical or real numbers, predicting the probability of an event occurring. The model is expressed as:

$$P(Y=1 \mid X) = \frac{e^{\beta 0 + \beta 1X1 + \beta 2X2 + \dots + \beta kXk}}{1 + e^{\beta 0 + \beta 1X1 + \beta 2X2 + \dots + \beta kXk}}$$

Where:

P(Y=1|X) is the probability that the dependent variable Y equals 1 given X.

 X_k are the independent variables, β_k are the coefficients to be estimated.

e is the base of the natural logarithm.

The model examines how independent variables influence the likelihood of specific economic outcomes, providing insights into the directional relationship between predictors and outcomes.

Transformative First Difference (FD) Method

The First Difference Method is used in regression analysis to address issues like non-stationarity and omitted variable bias by analysing changes between consecutive observations. It transforms the data by computing differences, making the model:

$$\Delta Y_t = \beta \Delta X_t + \Delta \epsilon_t$$

where $\Delta Y_t = Y_t - Y_{t-1}$ and $\Delta X_t = X_t - X_{t-1}$. This method eliminates time-invariant unobserved effects, focusing on the variation within the data. It is commonly applied in time-series and panel data analysis.

Variance Inflation Factor (VIF) – Multicollinearity Test

The Variance Inflation Factor (VIF) is used to detect multicollinearity in regression models by measuring how much the variance of a regression coefficient is inflated due to correlation with other predictors. The formula for VIF is:

 $\text{VIF}_i = \frac{1}{1 - R_i^2}$

where R_i^2 is the coefficient of determination obtained by regressing the i-th predictor on all other predictors. A high VIF (typically > 10) indicates significant multicollinearity, which may distort the regression results and reduce the reliability of the coefficients.

Adjusted R-squared

The Adjusted R-squared adjusts the R-squared value for the number of predictors in a regression model, providing a more accurate measure of goodness-of-fit, especially with multiple predictors. The formula is:

$$\bar{R}^2 = 1 - \frac{(1 - R^2)(n - 1)}{n - p - 1}$$

where R^2 is the R-squared value, n is the number of observations, and pp is the number of predictors. Unlike R-squared, the Adjusted R-squared penalizes unnecessary variables, preventing overfitting and giving a more reliable evaluation of model performance.

Standard Error (SE)

Standard Error (SE) measures the precision of a sample statistic, such as the mean, relative to the population parameter. It is calculated as:

$$SE = \frac{\sigma}{\sqrt{n}}$$

where σ is the population standard deviation and n is the sample size. A smaller SE indicates greater accuracy of the sample estimate, making it critical in hypothesis testing and confidence interval calculation.

Akaike Information Criterion (AIC)

The Akaike Information Criterion (AIC) is used to evaluate and compare the goodness of fit of statistical models, balancing model complexity and fit. The formula for AIC is:

AIC = 2k - 2ln(L)

where k is the number of parameters in the model, and L is the likelihood of the model. A lower AIC value indicates a better-fitting model, while penalizing excessive complexity. It is widely used in model selection, especially when comparing models with different numbers of parameters.

where $\Delta Y_t = Y_t - Y_{t-1}$ and $\Delta X_t = X_t - X_{t-1}$. This method eliminates time-invariant unobserved effects, focusing on the variation within the data. It is commonly applied in time-series and panel data analysis.

Durbin-Watson (DW) Test

The Durbin-Watson (DW) Test checks for autocorrelation in the residuals of a regression model, particularly for first-order correlation. The test statistic is:

$$\mathbf{DW} = \frac{\sum_{t=2}^{n} (\hat{\varepsilon}_t - \hat{\varepsilon}_{t-1})^2}{\sum_{t=1}^{n} \hat{\varepsilon}_t^2}$$

where $\hat{\varepsilon}_t$ are the residuals at time t. The DW statistic ranges from 0 to 4; a value near 2 indicates no autocorrelation, values < 2 suggest positive autocorrelation, and values > 2 indicate negative autocorrelation. This test is critical for ensuring the validity of regression assumptions in time-series data.

Log-likelihood

Log-likelihood quantifies how well a statistical model fits the observed data by calculating the logarithm of the likelihood function, which represents the probability of the observed outcomes given the model parameters. In regression, maximizing the log-likelihood helps identify parameter estimates that best explain the data. It is expressed as:

$$\ln(\mathbf{L}) = \sum_{i=1}^{n} \ln f(y_i \mid X_i, \beta)$$

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where $f(y_i | X_i, \beta)$ is the probability density or mass function, y_i are the observed values, X_i are the predictors, and represents the model parameters.

Normality test

The Chi-square test for normality is used to assess whether a dataset follows a normal distribution. It compares the observed frequency of data in each category with the expected frequency if the data were normally distributed. The formula for the Chi-square test is:

$$\chi^2 = \Sigma \frac{(O-E)^2}{E}$$

where O is the observed frequency, E is the expected frequency, and the summation is over all categories. A high Chi-square value indicates a significant deviation from normality.

Breusch-Pagan (BP) Test

The Breusch-Pagan (BP) Test detects heteroscedasticity in regression models by assessing whether error variances depend on independent variables. It involves regressing the squared residuals ($\hat{\varepsilon}^2$) on the predictors:

 $\hat{\varepsilon}^2 = \alpha_0 + \alpha_1 X_1 + \alpha_2 X_2 + \dots + \alpha_k X_k + u$

The test statistic is: BP = $\frac{1}{2}R_{aux}^2$ n

where R_{aux}^2 is the coefficient of determination from the auxiliary regression. The BP statistic follows a chi-squared distribution, with higher values indicating heteroscedasticity.

Lagrange Multiplier (LM) Test

The Lagrange Multiplier (LM) Test for autocorrelation detects serial correlation in residuals of a regression model. It involves regressing residuals ($\hat{\varepsilon}_t$) on lagged residuals and independent variables. The auxiliary regression is:

$$\hat{\varepsilon}_t = \alpha_0 + \alpha_1 \hat{\varepsilon}_{t-1} + \alpha_2 \hat{\varepsilon}_{t-2} + \dots + \alpha_p \hat{\varepsilon}_{t-p} + ut$$

The test statistic is: $LM = nR^2$

where n is the sample size, and R^2 is the auxiliary regression's determination coefficient. The LM statistic follows a chi-squared distribution, with significance indicating autocorrelation.

Lagrange Multiplier (LM) Test for ARCH Effect

The Lagrange Multiplier (LM) Test for ARCH Effect identifies autoregressive conditional heteroscedasticity (ARCH) in time-series data. It involves regressing squared residuals ($\hat{\varepsilon}_t$) on their lagged values. The auxiliary regression is:

$$\hat{\varepsilon}_t = \alpha_0 + \alpha_1 \hat{\varepsilon}_{t-1} + \alpha_2 \hat{\varepsilon}_{t-2} + \dots + \alpha_p \hat{\varepsilon}_{t-p} + \text{ut}$$

The test statistic is: $LM = nR^2$

where n is the sample size, and R^2 is from the auxiliary regression. A significant LM statistic indicates ARCH effects, essential for volatility modelling.

Brock-Dechert-Scheinkman (BDS) Test

The Brock-Dechert-Scheinkman (BDS) Test assesses non-linearity or dependence in time-series data by examining deviations from randomness. It compares the correlation of points in reconstructed phase space at varying dimensions. The test statistic is:

$$W = \frac{\sqrt{n} (C_m(\varepsilon) - C_1^m(\varepsilon))}{\sigma_m(\varepsilon)}$$

where $C_m(\varepsilon)$ is the correlation integral for dimension m, $C_1^m(\varepsilon)$ is the product of one-dimensional correlation integrals, and $\sigma_m(\varepsilon)$ is the standard deviation. A significant result indicates non-linear structure, making the test vital for analysing chaotic or complex systems.

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Table 1: Descriptive statistics of macro-economic variables.												
Variable	Ν	Mean	Median	Minimum	Maximum	Std. Dev.	C.V.	Skewness	Ex. kurtosis	IQ range	Jarque-Bera	
10Y GSec	72	6.9268	7.0453	5.8297	8.0157	0.6083	0.0878	-0.2084	-0.9974	0.9922	3.5064 (0.1732)	
91Day TB	72	5.2229	5.5964	3.0473	7.1443	1.4463	0.2769	-0.2699	-1.5833	3.0061	8.3945* (0.0150)	
Foreign Reserves	72	11.11	10.532	-8.5214	30.394	8.8492	0.7965	0.0969	-0.2498	10.905	0.2999 (0.8607)	
Interest Rate	72	5.5285	5.651	4.2531	6.7535	1.0353	0.1872	-0.1552	-1.6569	2.2562	8.5249* (0.0141)	
Inflation	72	5.2999	5.5315	1.971	7.7912	1.5097	0.2848	-0.3206	-0.7932	2.375	3.1215 (0.2099)	
Gold	72	7.366	6.2209	-9.1408	32.296	10.853	1.4734	0.3771	-0.9333	18.981	4.3195 (0.1153)	
NSE NIFTY	72	11.246	9.9375	-30.156	53.571	15.231	1.3544	0.5067	1.1754	13.151	7.2259* (0.0269)	
IIP	72	4.1495	3.9477	-57.312	133.52	18.939	4.5641	3.9241	30.026	5.3214	2889.5* (0.0000)	
FDI	72	-7.2366	1.0125	-477.04	368.43	111.43	15.398	-0.4208	4.6886	117.48	68.074* (0.0000)	
Exchange Rate	72	4.0137	3.4639	-5.7986	13.367	4.1957	1.0453	-0.0568	-0.6265	6.8366	1.2162 (0.5443)	

IV. **Result Analysis**

Source: The Authors, Note: *p < 0.05.

Table 2: Multivariate Correlations of the macro-economic variables.

Particulars	10Y GSec	91Day TB	Foreign Reserves	Interest Rate	Inflation	Gold	NSE NIFTY	IIP	FDI	Exchange Rate
10Y GSec	1	0.8527*	-0.6796*	0.8030*	-0.3084*	-0.5230*	-0.2184	0.0151	-0.0004	0.4460*
91Day TB	0.8527*	1	-0.6232*	0.9872*	-0.4160*	-0.2830	-0.3091*	-0.0514	-0.0907	0.3540*
Foreign Reserves	-0.6796*	-0.6232*	1	-0.5531*	0.1585	0.6022*	0.0391	-0.0798	0.0556	-0.2990
Interest Rate	0.8030*	0.9872*	-0.5531*	1	-0.4512*	-0.2457	-0.3040*	-0.0461	-0.1111	0.3521*
Inflation	-0.3084*	-0.4160*	0.1585	-0.4512*	1	0.3127*	-0.0892	-0.1235	-0.0658	-0.0248
Gold	-0.5230*	-0.2830	0.6022*	-0.2457	0.3127*	1	-0.4254*	-0.3065*	-0.0659	-0.2415
NSE NIFTY	-0.2184	-0.3091*	0.0391	-0.3040*	-0.0892	-0.4254*	1	0.5277*	0.2321	-0.6109*
IIP	0.0151	-0.0514	-0.0798	-0.0461	-0.1235	-0.3065*	0.5277*	1	0.2951	-0.2832
FDI	-0.0004	-0.0907	0.0556	-0.1111	-0.0658	-0.0659	0.2321	0.2951	1	-0.2302
Exchange Rate	0.4460*	0.3540*	-0.2990	0.3521*	-0.0248	-0.2415	-0.6109*	-0.2832	-0.2302	1

Source: The Authors, Note: *p < 0.05.

The descriptive statistics in Table 1 provide critical insights into the behaviour of key macroeconomic variables across the study period. The 10-year Government Securities (10Y GSec) yield exhibits a relatively narrow range, with a mean of 6.9268% and a standard deviation of 0.6083, indicating low variability. Skewness (-0.2084) and excess kurtosis (-0.9974) suggest a slightly left-skewed distribution, and the Jarque-Bera test confirms normality (p > 0.05). In contrast, the 91-day Treasury Bill (TB) rate shows a higher standard deviation of 1.4463, reflecting greater volatility, with a significant Jarque-Bera result (p < 0.05), indicating non-normality in its distribution.

Foreign reserves exhibit a wide range, from -8.5214 to 30.394, with an average value of 11.11 and relatively high variability (standard deviation of 8.8492). The normal distribution assumption for these variable holds, as indicated by the Jarque-Bera test (p > 0.05). The Interest rate shows moderate variability (standard deviation 1.0353), with a near-normal distribution except for slight deviations, supported by a significant Jarque-Bera value (p < 0.05). Inflation has a mean of 5.2999% with relatively low variability (standard deviation 1.5097), and the Jarque-Bera test suggests normality (p > 0.05). Gold returns demonstrate substantial variability, with a high standard deviation of 10.853 and a significant skewness of 0.3771, indicating a non-normal distribution confirmed by the Jarque-Bera value.

Equity market performance, represented by NSE NIFTY, exhibits the highest volatility among the variables, with a standard deviation of 15.231 and a substantial range from -30.156 to 53.571. Skewness (0.5067) and excess kurtosis (1.1754) suggest a right-skewed and leptokurtic distribution, corroborated by a significant Jarque-Bera test (p < 0.05). Industrial production (IIP) displays extreme values, with a standard deviation of 18.939 and positive skewness (3.9241), indicating high variability. A significant Jarque-Bera test (p < 0.05) reflects severe deviations from normality. Foreign direct investment (FDI) has the most substantial dispersion, with a standard deviation of 111.43, and its non-normality is evident from significant kurtosis (4.6886) and the Jarque-Bera test (p < 0.05).

When analysed parameter-wise, central tendency measures like mean and median indicate that most variables, except FDI, exhibit relatively close central values, hinting at stable trends. Variability, measured through standard deviation and coefficient of variation (C.V.), is pronounced in financial market indicators like NSE NIFTY and Gold, suggesting susceptibility to external shocks. Skewness and kurtosis values reveal asymmetries and tail behaviour, with variables like IIP and FDI showing extreme departures from normality, underscoring their volatile nature.

The multivariate correlation analysis in Table 2 highlights significant relationships between key macroeconomic variables. The strong positive correlation (0.8527*) between the 10-year Government Securities (10Y GSec) yield and the 91-day Treasury Bill (TB) rate indicates a synchronized movement between long-term and short-term bond yields. Additionally, the 10Y GSec shows a moderate positive correlation with the Interest rate (0.8030*), suggesting a direct relationship with monetary policy decisions. However, the inverse correlation with foreign reserves (-0.6796*) and gold (-0.5230*) highlights the impact of external and commodity market factors on long-term yields.

The 91-day TB rate is similarly influenced by the Interest rate, with a very strong positive correlation (0.9872*), underlining its dependence on short-term monetary policy. It also shows significant inverse relationships with foreign reserves (-0.6232*) and inflation (-0.4160*), reflecting the role of inflationary pressures and external stability in shaping short-term yields. Interestingly, while the TB rate correlates positively with the exchange rate (0.3540*), its negative association with NSE NIFTY (-0.3091*) suggests a divergence between fixed-income securities and equity market performance.

Foreign reserves show a moderate positive correlation with gold (0.6022*), highlighting their complementary role as safety assets during economic instability. Conversely, reserves exhibit a negative relationship with key monetary policy variables, such as the 10Y GSec (-0.6796*), 91-day TB (-0.6232*), and Interest rate (-0.5531*), emphasizing the counter-cyclicality of reserve accumulation. Interestingly, foreign reserves show weak and non-significant correlations with inflation (0.1585), IIP (-0.0798), and FDI (0.0556), indicating a limited direct impact of these factors on reserve levels.

Gold exhibits a complex relationship with other variables. It correlates positively with inflation (0.3127*), which is expected as gold often serves as an inflation hedge. However, its negative correlations with the 10Y GSec (-0.5230*), Interest rate (-0.2457), and NSE NIFTY (-0.4254*) suggest an inverse relationship with traditional financial assets, underscoring its role as a safe-haven investment during periods of economic uncertainty. Additionally, gold's negative correlation with IIP (-0.3065*) highlights its counter-cyclical behaviour during industrial slowdowns.

Equity market performance, represented by NSE NIFTY, shows a significant positive correlation with IIP (0.5277*), reflecting the influence of industrial growth on equity returns. However, its negative correlation with gold (-0.4254*) and the exchange rate (-0.6109*) implies that currency depreciation and rising gold prices often coincide with lower equity performance. The weak and negative correlations between NSE NIFTY and bond yields, as well as monetary policy variables, suggest limited direct interaction between equity and fixed-income markets in the short term.

Finally, the exchange rate exhibits a positive correlation with short-term bond yields (91-day TB at 0.3540*) and the Interest rate (0.3521*), suggesting that currency movements are sensitive to changes in monetary policy. However, its negative correlation with NSE NIFTY (-0.6109*) and foreign reserves (-0.2990) implies that depreciation often coincides with lower equity performance and reserve depletion. The relatively weak correlations between the exchange rate and inflation (-0.0248) or FDI (-0.2302) highlight limited short-term interactions between these variables. Overall, these findings underscore the intricate interplay of monetary, financial, and external factors in shaping macroeconomic outcomes.

The multivariate correlation analysis reveals notable differences in the relationships among macroeconomic variables, highlighting the unique interplay between monetary policy, financial markets, and external factors. Long-term (10Y GSec) and short-term (91-day TB) yields show strong alignment, reflecting consistent monetary policy influence, while equity markets (NSE NIFTY) and safe-haven assets like gold exhibit inverse relationships, underscoring their contrasting responses to economic uncertainty. External stability indicators, such as foreign reserves and the exchange rate, correlate negatively with bond yields and equity returns, emphasizing their counter-cyclicality during volatile periods. Industrial growth (IIP) aligns positively with equity returns, showcasing its crucial role in driving market performance. These results highlight the divergent behaviour of financial instruments and economic variables, offering critical insights into their interdependencies in different market conditions.

The analysis of Table 3 focuses on the variables influencing the 10-year Government Security (10Y GSec) yields across different regression models—OLS, HSC, Tobit, Logistic, and FD. These models provide varied perspectives on how macroeconomic variables and financial indicators impact long-term bond yields. NSE NIFTY, representing equity market returns, exhibits a negative and statistically significant relationship with 10Y GSec yields in the OLS, HSC, and Logistic regressions. This relationship underscores the inverse link between equity market performance and bond yields, as strong equity markets typically draw investments away from bonds, reducing their demand and increasing yields. The Interest Rate, a key monetary policy tool, demonstrates a robust and positive influence on bond yields across OLS, HSC, and Logistic regressions, with high significance levels. The consistent impact reflects the central role of the Interest Rate in determining borrowing costs and influencing investor expectations about future interest rates. Similarly, Inflation has a positive and significant impact, reinforcing its role in shaping long-term bond yields through inflation expectations and risk premiums. These findings align with conventional financial theories where higher inflation typically leads to higher yields as investors demand compensation for erosion in purchasing power.

The impact of Gold prices on 10Y GSec yields is consistently negative across OLS and HSC models, indicating its safe-haven status. Rising gold prices suggest heightened risk aversion, reducing the demand for risky assets, including long-term bonds. This relationship, however, is weaker in the Tobit and FD regressions, possibly due to the latter models' focus on different assumptions about the dependent variable distribution. FDI and Exchange Rate exhibit mixed significance, with their effects being model-dependent. While FDI shows marginal significance in OLS and HSC, its impact diminishes in Logistic and FD regressions. Exchange Rate, although insignificant in OLS, gains prominence in the Tobit and HSC models, suggesting that currency stability indirectly affects bond market dynamics. When comparing model fit, the HSC regression emerges as the most robust model with the highest Adjusted R-squared (0.885), indicating it captures the greatest variance among the variables. The Logistic regression stands out with the lowest Akaike Information Criterion (AIC) of -241.2830 and the highest Log-Likelihood (129.6415), indicating an excellent fit for the data. The OLS model, while straightforward, exhibits moderate performance with an Adjusted R-squared of 0.794 and AIC of 27.0925. Conversely, the FD regression performs poorly, with low significance levels across variables and high AIC, indicating limited explanatory power.

Further statistical tests provide additional insights. The Durbin-Watson statistic highlights mild autocorrelation in the residuals of OLS, HSC, and Logistic regressions. The Breusch-Pagan test for heteroscedasticity is insignificant in OLS and Logistic regressions but significant in FD regression, raising concerns about the reliability of the latter. Additionally, the Chow test and QLR test for structural breaks confirm significant regime changes before and after COVID-19, emphasizing the temporal nature of bond yield determinants. In terms of residual error and overall reliability, the HSC and Logistic regressions outperform others. The Logistic regression achieves the best trade-off between bias and variance, as indicated by its lower standard error (S.E.) of regression and superior likelihood-based criteria. The HSC model, with its high Adjusted R-squared and significant coefficients, also effectively explains the variability in bond yields while accounting for potential heteroscedasticity in the data. On the other hand, the Tobit regression, while effective in addressing censored observations, fails to achieve the best fit compared to HSC and Logistic models, as evidenced by its higher AIC and lower log-likelihood. HSC and Logistic regressions provide the most reliable insights into the drivers of 10Y GSec yields, outperforming simpler models like OLS and specialized models like Tobit and FD.

Table 3: Variables impact on 10Y Government Security yields and comparison of different regression models												
Variables and Pasiduals Test	OLS	S Regression		HSC Reg	ression	Tobit Regr	ression	Logistic Re	gression	FD Reg	ression	
variables and Residuals Test	Coefficient	p-value	VIF	Coefficient	p-value	Coefficient	p-value	Coefficient	p-value	Coefficient	p-value	
Constant	4.88819***	< 0.0001		4.8629***	< 0.0001	4.12586***	< 0.0001	-2.9179***	< 0.0001	0.00203	0.9403	
NSE NIFTY	-0.01158**	0.0202	5.120	-0.0107***	0.0052	0.00354	0.4010	-0.00184**	0.0173	-0.00455	0.2757	
Interest Rate	0.39123***	< 0.0001	2.234	0.39408***	< 0.0001	0.45297***	< 0.0001	0.06127***	< 0.0001	0.21380	0.1893	
Inflation	0.05547**	0.0451	1.570	0.06244***	0.0025	0.11475***	< 0.0001	0.00907**	0.0350	0.06518	0.0812	
IIP	0.00019	0.9287	1.491	0.00006	0.9650	0.00118	0.3680	0.00002	0.9469	0.00059	0.6940	
Gold	-0.02858***	< 0.0001	4.785	-0.0321***	< 0.0001	-0.00630886	0.1866	-0.0044***	< 0.0001	-0.00625	0.3460	
FDI	0.00055*	0.0857	1.138	0.00057**	0.0195	0.00001	0.9671	0.00008*	0.0867	0.00015	0.5235	
Exchange Rate	-0.01059	0.4727	3.539	-0.01807**	0.04080	0.04396**	0.0131	-0.00212	0.3560	-0.00897	0.6179	
Foreign Reserves	-0.00286	0.6383	2.693	-0.00144	0.74710	-0.03208***	< 0.0001	-0.00076	0.4203	-0.00621	0.6233	
S.E. of Regression	0.275538			1.673415		-	-		0.042734 ^a		0.226383	
Adjusted R-squared	0.794874			0.8851	0.885156 ^b		-		0.800587		0.01391	
Akaike Criterion (AIC)		27.0925		286.8	537	0.0515	87	-241.2830ª		-1.079380		
Log-likelihood	-	4.546251		-134.4268		9.9742	.06	129.6415 ^b		9.53969		
Durbin-Watson	().927634		0.903055		-	-		0.906624		2.134965	
F Stat	35.391	16*** (0.0000)		69.40365*** (0.0000)		-		36.63057*** (0.0000)		0.879986 (0.5383)		
Chi-square		-		-	-		323.5763*** (0.0000)		-		-	
Sigma		-		-		0.161754** (0.0197)		-		-		
Left-censored observations		-		-		0		-		-		
Right-censored observations		-		-		38		-		-		
Normality (Chi-square)	10.25	4*** (0.0059)		7.51215**	(0.0233)	8.5748** (0	0.0137)	2.83285 (0.2425)	3.57556	(0.1673)	
Non-linearity test (Chi-square)	1.87	966 (0.3906)		-		-		-		7.86867	(0.4464)	
Breusch-Pagan test for HS (LM)	10.0	587 (0.2609)		-		-		-		17.2327**	^c (0.0277)	
Autocorrelation (LMF)	2.7609	91*** (0.0057)		-		-		-		0.849198	(0.6012)	
ARCH (LM)	13.2	197 (0.3532)		-		-		-		2.31816	(0.1278)	
QLR test for structural break	81.780	04*** (0.0000)		-		-	-		-		^c (0.0429)	
Chow Test structural break	6.2727	75*** (0.0000)								2.8669***	^c (0.0078)	
RESET test specification	0.844	4503 (0.4347)		-		-		-		3.92828** (0.0249)		

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Source: The Authors. Note: ***p < 0.01, **p < 0.05 & *p < 0.10, ^a Lowest value, and ^b Highest Value.

Table 4: Variables impact on 91Day Treasury Bills yields and comparison of different regression models												
Variables and Pesiduals Test	OLS	S Regression		HSC Regr	ession	Tobit Reg	ression	Logistic Re	gression	FD Regr	ession	
variables and Residuals Test	Coefficient	p-value	VIF	Coefficient	p-value	Coefficient	p-value	Coefficient	p-value	Coefficient	p-value	
Constant	-1.7255***	< 0.0001	-	-2.1176***	< 0.0001	-1.7173***	< 0.0001	-4.3554***	< 0.0001	0.00594	0.8115	
NSE NIFTY	-0.00791**	0.0163	5.120	-0.0081***	0.0002	-0.0079***	0.0073	-0.0033***	0.0002	-0.00529	0.1696	
Interest Rate	1.30445***	< 0.0001	2.234	1.36916***	< 0.0001	1.30337***	< 0.0001	0.2741***	< 0.0001	0.6395***	< 0.0001	
Inflation	0.02864	0.1149	1.570	0.02431*	0.0681	0.02903*	0.0800	0.00918**	0.0479	-0.01328	0.6961	
IIP	-0.00093	0.5086	1.491	-0.00048	0.5515	-0.00094	0.4676	-0.00034	0.3453	-0.00023	0.8675	
Gold	-0.00793*	0.0728	4.785	-0.01121***	< 0.0001	-0.00788*	0.0502	-0.00262**	0.0211	-0.00202	0.7402	
FDI	0.00031	0.1334	1.138	0.00037**	0.0352	0.00030	0.1192	0.00009*	0.0787	0.00022	0.2993	
Exchange Rate	-0.02254**	0.0231	3.539	-0.02889***	< 0.0001	-0.0233***	0.0096	-0.0094***	0.0003	-0.00586	0.7228	
Foreign Reserves	-0.01538*** 0.0003 2.693		-0.00906**	0.01200	-0.0156***	< 0.0001	-0.0047***	< 0.0001	-0.02156	0.0672		
S.E. of Regression	0.181898			1.5032	04	-		0.04622ª		0.208422		
Adjusted R-squared	0.984181			0.9946	69 ^b	-		0.97851		0.262484		
Akaike Criterion (AIC)	-	-32.70733		271.40	71	-29.04	919	-229.9940ª		-12.81773		
Log-likelihood		25.35366		-126.70	035	24.52	46	123.997 ^b		15.40	887	
Durbin-Watson		1.795393		1.625213		-		1.47655		2.234	251	
F Stat	553.16	97*** (0.0000)		1656.993*** (0.0000)		-		405.1122*** (0.0000)		4.11415*** (0.0005)		
Chi-square		-		-		5146.959*** (0.0000)		-		-		
Sigma		-		-		0.168332** (0.0141)		-		-		
Left-censored observations		-		-		0		-		-		
Right-censored observations		-		-		1		-		-		
Normality (Chi-square)	5.59	254 (0.0610)		16.6338***	(0.0002)	8.47729**	(0.0144)	0.807158 (0.6679)	30.09*** ((0.0000)	
Non-linearity test (Chi-square)	21.850	65*** (0.0051)		-		-		-		10.703 (0).2191)	
Breusch-Pagan test for HS (LM)	53.4	376 (0.1557)		-		-		-		56.1751 (0.1031)	
Autocorrelation (LMF)	0.640	0719 (0.7974)		-		-		-		1.46845 (0.1679)	
ARCH (LM)	13.1	572 (0.3577)		-		-		-		4.20991 (0.9793)	
QLR test for structural break	28.92	73** (0.0154)		-		-		-		22.0361 (0.1309)	
Chow Test structural break	2.029	46** (0.0484)								1.2306 (0).2967)	
RESET test specification	1.65	783 (0.1990)		-		-		-		0.0436961	0.0436961 (0.9572)	

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Source: The Authors. Note: ***p < 0.01, **p < 0.05 & *p < 0.10, ^a Lowest value, and ^b Highest Value.

Table 4 presents a comprehensive analysis of the determinants of 91-Day Treasury Bill (TB) yields, exploring the influence of macroeconomic variables through five regression models: OLS, HSC, Tobit, Logistic, and FD. Among these, the NSE NIFTY index, which represents equity market performance, consistently shows a negative relationship with TB yields. This indicates that when equity markets perform well, investors tend to shift their investments from Treasury Bills to stocks, reducing demand for short-term debt and thus raising yields. Conversely, during periods of market uncertainty, demand for safer assets like Treasury Bills increases, pushing yields lower.

The Interest Rate stands out as the most influential determinant of 91-Day TB yields, demonstrating a strong positive relationship across all regression models. This highlights the central role of monetary policy in shaping short-term borrowing costs. A higher interest rate typically signals a tightening of monetary policy, which increases the cost of borrowing and raises yields on Treasury Bills as investors seek compensation for potential inflation and higher rates. The consistent statistical significance of this relationship across all models emphasizes its key role in determining short-term debt instrument yields.

Inflation also influences short-term Treasury yields, although its impact is not as consistent. In the OLS, HSC, and Tobit models, inflation has a marginally significant positive effect on TB yields, suggesting that rising inflation leads to higher yields as investors demand compensation for the erosion of purchasing power. However, the FD regression model shows no significant relationship between inflation and TB yields, indicating that this model may not capture the short-term variations in inflation adequately, thus highlighting its limitations in modelling inflation's effect on short-term debt instruments.

Gold prices exhibit a negative relationship with 91-Day TB yields in most models, indicating that rising gold prices, often a sign of increased market risk or uncertainty, lead to lower demand for Treasury Bills. As investors seek safer assets like gold, demand for short-term bonds decreases, pushing yields down. This relationship is most evident in the OLS, HSC, and Tobit regressions, where gold's negative impact is statistically significant. However, in the FD model, the relationship weakens, suggesting that the FD regression may not adequately capture the risk-averse behaviour of investors in response to rising gold prices.

Foreign Reserves have a consistent negative impact on 91-Day TB yields across most regression models. This relationship suggests that higher levels of foreign reserves help stabilize the domestic economy and reduce the need for high short-term yields. By bolstering investor confidence and mitigating external shocks, foreign reserves reduce the risks associated with Treasury Bills, leading to lower yields. The consistent negative influence of foreign reserves highlights their role in enhancing market stability and influencing short-term borrowing costs. In terms of model performance, the HSC regression emerges as the most robust, achieving the highest Adjusted R-squared (0.9947), indicating its strong explanatory power. This model captures the largest proportion of variance in 91-Day TB yields, outperforming others in terms of fit. The Logistic regression follows closely, demonstrating high efficiency with the lowest AIC and highest Log-Likelihood, indicating an excellent model fit. While the OLS model offers simplicity and good explanatory power, it lags behind in precision. The FD regression, with weak performance across the fit metrics, is the least effective for analysing 91-Day TB yields, suggesting its limited usefulness in this context.

Table 5 focuses on the analysis of inflation determinants across five regression models: OLS, HSC, Tobit, Logistic, and FD. The Interest Rate consistently emerges as the most significant determinant of inflation across all models. A negative and statistically significant relationship is observed in the OLS, HSC, Tobit, and Logistic regressions, indicating that higher interest rates typically reduce inflationary pressures. Tightening monetary policy, often associated with rising interest rates, is a primary tool for central banks to control inflation, and this relationship is evident in the results. However, the FD regression shows a less significant relationship, highlighting the limitations of this model in capturing the dynamic impact of interest rates on inflation. The 10-Year Government Securities (10Y GSec) yield shows a positive relationship with inflation across all models, suggesting that long-term bond yields are sensitive to inflation expectations. The significant positive coefficients in the OLS, HSC, Tobit, and Logistic regressions point to the role of long-term borrowing costs in reflecting inflation expectations. As inflation rises, investors demand higher yields on long-term debt to compensate for the anticipated erosion in purchasing power. This relationship reinforces the conventional economic theory that higher inflation results in higher bond yields as investors seek compensation for future inflation risks.

Gold prices, often viewed as a hedge against inflation, consistently show a positive and statistically significant relationship with inflation across the OLS, HSC, Tobit, and Logistic models. The strong positive coefficients underscore gold's role as a store of value during periods of inflationary pressure. As inflation rises, investors tend to move toward gold to protect their wealth from the eroding purchasing power of fiat currencies. This relationship is particularly significant in the OLS and HSC regressions, highlighting gold's status as an important inflation hedge. However, the FD model fails to capture this relationship, indicating its limitations in addressing the broader impact of gold on inflation expectations. Foreign Reserves demonstrate a consistent negative relationship with inflation in the OLS, HSC, Tobit, and Logistic regressions. The negative coefficients suggest that higher foreign reserves help stabilize the economy and reduce inflationary pressures. Reserves can

mitigate external shocks and support the domestic currency, thus lowering inflation expectations. This relationship aligns with the notion that countries with higher reserves are better positioned to manage external imbalances, which can contribute to maintaining stable inflation rates. However, similar to other variables, the FD regression fails to detect a meaningful relationship, reflecting its limitations in capturing short-term fluctuations in inflation.

The fit metrics reveal notable differences in model performance. The HSC regression stands out with the highest Adjusted R-squared (0.5615), indicating it captures the largest amount of variance in inflation. This suggests that the HSC model provides the most comprehensive view of inflation determinants. The Logistic regression follows closely with a strong model fit, evidenced by its lowest AIC and highest Log-Likelihood. While the OLS regression offers simplicity and relatively strong results, it lags behind in terms of model fit, with a lower Adjusted R-squared and less precise coefficients. The FD regression, on the other hand, exhibits the lowest Adjusted R-squared and weaker significance levels across key variables, highlighting its limited ability to explain inflation dynamics. Overall, Table 5 highlights the significant relationships between macroeconomic variables and inflation, with the Interest Rate, 10Y GSec yield, Gold prices, and Foreign Reserves emerging as the most influential factors. The HSC and Logistic regressions provide the most reliable insights into these relationships, outperforming the simpler OLS model and the more specialized FD regression.

Table 6 presents the analysis of factors influencing interest rates across five regression models: OLS, HSC, Tobit, Logistic, and FD regressions. It compares variable coefficients, significance levels, and the fit of the models using metrics such as Adjusted R-squared, Standard Error (SE), Log-Likelihood, and Akaike Criterion (AIC). This analysis reveals the relative performance of each model in explaining interest rate variability. The 10Y GSec (10-Year Government Security yield) emerges as the most consistent determinant of interest rates. It shows a strong positive and highly significant relationship across OLS, HSC, Tobit, and Logistic regressions (p < 0.01). This finding aligns with economic theory, as long-term bond yields reflect broader interest rate trends. However, in the FD regression, the variable's coefficient is smaller and statistically insignificant, suggesting the model's limitation in capturing this relationship effectively.

Inflation exhibits a significant and negative relationship with interest rates across all models, with the strongest significance observed in OLS, HSC, Tobit, and Logistic regressions (p < 0.01). This inverse relationship highlights the role of inflation expectations in influencing real interest rates. While the FD regression maintains a negative coefficient, the significance drops to p < 0.1, indicating weaker explanatory power in this model. Gold prices, a proxy for market expectations and economic uncertainty, have a consistently positive and significant impact on interest rates in OLS, HSC, Tobit, and Logistic regressions (p < 0.01). These finding underscores gold's role as a hedge against inflation, which influences monetary policy and interest rates. However, the FD model fails to capture this relationship, with an insignificant coefficient.

Foreign Reserves negatively influence interest rates in OLS, HSC, Tobit, and Logistic regressions, with significance levels varying between p < 0.05 and p < 0.01. This relationship suggests that higher reserves reduce external borrowing pressures and stabilize domestic interest rates. As with other key variables, the FD regression fails to detect a significant relationship. In terms of model fit, the HSC regression demonstrates the highest Adjusted R-squared (0.9201), indicating its superior explanatory power compared to the other models. It also achieves significant coefficients for most variables, making it the most robust model. The Logistic regression stands out for its lowest AIC (-121.1925) and highest Log-Likelihood (69.59625), showing its efficiency in modelling the data despite its focus on classification over continuous predictions.

The Tobit regression performs well for censored data, maintaining significant coefficients for most variables. However, its Adjusted R-squared is not directly comparable to OLS or HSC. The FD regression, with the lowest Adjusted R-squared (0.0619), fails to provide meaningful insights and performs poorly across fit metrics. OLS regression offers simplicity and retains high significance levels but lags behind HSC and Logistic models in terms of overall model fit. Overall, the HSC regression is the most comprehensive and effective model for analysing the determinants of interest rates, followed by the Logistic regression for its efficiency and robust fit metrics. The Tobit model is valuable for censored datasets, while the FD regression struggles to capture significant relationships provide strong model fit. or а

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Variables and Residuals Test	0	LS Regression		HSC Reg	HSC Regression		gression	Logistic Re	egression	FD Regr	ession	
variables and residuals rest	Coefficient	p-value	VIF	Coefficient	p-value	Coefficient	p-value	Coefficient	p-value	Coefficient	p-value	
Constant	3.73524	0.2568	-	4.76571*	0.0760	2.62052	0.4547	-3.3980***	< 0.0001	0.01440	0.8774	
Interest Rate	-1.26929***	< 0.0001	3.401	-1.2157***	< 0.0001	-1.5523***	< 0.0001	-0.2798***	< 0.0001	-0.72081	0.2547	
10Y GSec	1.12126**	0.0451	5.152	0.98227**	0.0187	1.52159**	0.0127	0.26292**	0.0417	-0.18652	0.6961	
NSE NIFTY	0.02948	0.1949	5.433	0.01380	0.4664	0.02613	0.2719	0.00765	0.1455	0.01002	0.4908	
Gold	0.10586*** 0.0011 5.244		0.1049***	0.0002	0.11467***	0.0005	0.02378***	0.0015	0.01449	0.5245		
Exchange Rate	0.10986* 0.0946 3.412		0.06313	0.2470	0.11195*	0.0973	0.02056	0.1730	0.02293	0.7115		
FDI	-0.00116	0.4240	1.181	-0.00183	0.1642	-0.00147	0.3294	-0.00037	0.2659	-0.00011	0.8943	
IIP	-0.00109	0.9084	1.491	0.00935	0.3615	-0.00088	0.9305	-0.00022	0.9189	-0.01166**	0.0210	
Foreign Reserves	-0.06669**	0.0127	2.446	-0.0701***	0.0019	-0.0721***	0.0094	-0.01369**	0.0256	-0.02581	0.5636	
S.E. of Regression	1.23877			1.641374		-		0.285574^{a}		0.781102		
Adjusted R-squared	0.326728			0.5614	473 ^b	-	-		0.305779		0.071734	
Akaike Criterion (AIC)	243.546			284.0	697	234.6	746	32.24436 ^a		174.78	845	
Log-likelihood	-112.7730			-133.0)349	-107.2	3373	-7.122178 ^b		-78.39	0223	
Durbin-Watson		0.488115		0.555	0.555588		-		0.417272		067	
F Stat	5.30	6903*** (0.0000)	I	12.3632*** (0.0000)		-		4.909118***(0.0000)		1.676178 ((0.1222)	
Chi-square		-		-		43.0103*** (0.0000)		-		-		
Sigma		-		-		1.28254 (0.1288)		-		-		
Left-censored observations		-		-		6		-		-		
Right-censored observations		-		-		11	l	-		-		
Normality (Chi-square)	2.	31555 (0.3141)		2.83242 (0.2426)	19.4805***	* (0.0000)	5.88486*	(0.0527)	25.0071***	(0.0000)	
Non-linearity test (Chi-square)	32.5	5336*** (0.0000)		-		-		-		7.70816 (0	0.4624)	
Breusch-Pagan test for HS (LM)	9.	81789 (0.2780)		-		-		-		35.4121 (0	0.8187)	
Autocorrelation (LMF)	12.3	3241*** (0.0000)		-		-		-		5.36701***	(0.0000)	
ARCH (LM)	21.	6089** (0.0421)		-		-		-		7.49851 (0	0.8229)	
QLR test for structural break	138	.571*** (0.0000)		-		-		-		28.8291**	(0.0158)	
Chow Test structural break	15.3	3968*** (0.0000)		-		-		-	-		0.962337 (0.4811)	
RESET test for specification	4.2	5301** (0.0186)		-		-		-		0.406712 (0.6676)		

Explo	oring i	the L	Dynamics	of In	nflation,	Interest	Rates.	and Bond	Yields: A	Comprehe	ensive .
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 Table 5: Variables impact on Inflation and comparison of different regression models

Source: The Authors. Note: ***p < 0.01, **p < 0.05 & *p < 0.10, ^aLowest value, and ^b Highest Value.

Table 6: Variables impact on Interest Rates and comparison of different regression models												
Variables and Residuals Test	OI	S Regression		HSC Reg	gression	Tobit Reg	gression	Logistic R	egression	FD Reg	ression	
variables and residuals rest	Coefficient	p-value	VIF	Coefficient	p-value	Coefficient	p-value	Coefficient	p-value	Coefficient	p-value	
Constant	-2.81331**	0.0352		-0.94642	0.2137	-2.8133**	0.0214	-4.481***	< 0.0001	0.00581	0.7820	
10Y GSec	1.33300***	< 0.0001	2.629	1.0654***	< 0.0001	1.3330***	< 0.0001	0.2617***	< 0.0001	0.12923	0.1893	
Inflation	-0.21396***	< 0.0001	1.219	-0.1897***	< 0.0001	-0.21***	< 0.0001	-0.041***	< 0.0001	-0.05107*	0.0787	
NSE NIFTY	0.00805	0.3906	5.516	0.00128	0.8409	0.00805	0.3554	0.00096	0.5942	-0.00133	0.6842	
Exchange Rate	0.02433	0.3710	3.523	0.03102**	0.0437	0.02433	0.3354	0.00295	0.5743	0.01638	0.2391	
Gold	0.04490***	0.0007	5.179	0.03221***	0.0003	0.0449***	0.0001	0.0085***	0.0009	-0.00476	0.3554	
IIP	0.00137	0.7251	1.489	0.00421	0.3731	0.00137	0.7057	0.00024	0.7477	-0.00037	0.7459	
FDI	-0.00093	0.1136	1.147	-0.00022	0.6239	-0.00093*	0.0863	-0.00017	0.1278	0.00013	0.4560	
Foreign Reserves	-0.02601**	0.0181	2.471	-0.0277***	0.0001	-0.026***	0.0095	-0.00543*	0.0111	-0.00731	0.4560	
S.E. of Regression	0.508602			1.517471		-	-		0.098393 ^a		003	
Adjusted R-squared	0.758657			0.920	136 ^b	-		0.767971		0.061904		
Akaike Criterion (AIC)		115.356		272.7	673	117.3	117.356		-121.1925 ^a		-36.82473	
Log-likelihood		-48.67800		-127.3837		-48.6	7800	69.59625 ^b		27.41	237	
Durbin-Watson		0.607474		0.390	0118	-		0.62747		1.697	586	
F Stat	28.89	838*** (0.0000)		103.2513*** (0.0000)		-		30.3745*** (0.0000)		1.577405 (0.1500)		
Chi-square		-		-		264.214*** (0.0000)		-		-		
Sigma		-		-	-		0.47575** (0.0396)		-		-	
Left-censored observations		-		-		0	0			-		
Right-censored observations		-		-	-		0		-		-	
Normality (Chi-square)	2.6	3743 (0.2674)		1.59322 ((0.4508)	23.2016***	* (0.0000)	1.81648	(0.4032)	39.295***	(0.0000)	
Non-linearity test (Chi-square)	21.72	214*** (0.0054)		-		-		-		7.90611 (0.4426)	
Breusch-Pagan test for HS (LM)	17.8	027** (0.0227)		-		-		-		41.818***	(0.0000)	
Autocorrelation (LMF)	5.943	324*** (0.0000)		-		-		-		2.06848 (2	2.06848)	
ARCH (LM)	28.8	33*** (0.0041)		-		-		-		5.95044 ((0.9185)	
QLR test for structural break	93.5	44*** (0.0000)		-		-		-		46.803***	(0.0000)	
Chow Test structural break	2.21	248** (0.0351)		-		-		-	-		(0.0054)	
RESET test for specification	7.940	061*** (0.0008)		-		-		-		0.109172 (0.8967)		

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Source: The Authors. Note: ***p < 0.01, **p < 0.05 & *p < 0.10, ^aLowest value, and ^b Highest Value.

V. Conclusion

In conclusion, the multivariate analysis across multiple regression models has provided deep insights into the factors influencing long-term bond yields (10Y GSec), short-term Treasury Bill yields (91-day TB), inflation, and interest rates. The study found that key macroeconomic variables such as the Interest Rate, Inflation, Gold, and the 10Y GSec yield have significant and consistent relationships with bond yields across various models, highlighting the interplay between monetary policy and investor sentiment. Specifically, the positive relationship between the 10Y GSec yield and the Interest Rate across OLS, HSC, and Logistic models underlines the critical role of central bank actions in shaping long-term borrowing costs. Furthermore, the negative relationship between bond yields and equity market performance (NSE NIFTY) confirms the inverse dynamics of investment flows between riskier equity markets and safer fixed-income assets, particularly during periods of market uncertainty.

The analysis also illuminated the role of external factors, such as Gold and Foreign Reserves, in shaping macroeconomic outcomes. Gold prices were found to have a significant negative impact on bond yields, reflecting its status as a safe-haven asset during times of market stress, which diminishes demand for long-term bonds. Similarly, Foreign Reserves demonstrated a stabilizing effect on both inflation and interest rates, suggesting that countries with higher reserve levels are better able to withstand external shocks and reduce inflationary pressures. The observed relationships between these variables underscore the importance of both domestic economic policies and global economic conditions in influencing macroeconomic outcomes.

The comparison of different regression models revealed the superior performance of the HSC and Logistic models in explaining macroeconomic variability. The HSC regression emerged as the most robust model, consistently capturing a high percentage of variance across all dependent variables, while the Logistic regression excelled in terms of model efficiency and fit. The OLS model, while providing a reliable baseline, was outperformed by HSC and Logistic models in terms of explanatory power and fit metrics. The Tobit and FD regressions, while useful for specific contexts, exhibited lower performance, especially in capturing significant relationships and explaining the variability in the data. This indicates the importance of selecting the right model based on the characteristics of the data and the underlying relationships being studied.

Overall, the findings highlight the complex interdependencies between monetary policy, financial markets, and external factors in shaping macroeconomic dynamics. The study emphasizes that long-term bond yields are influenced not only by domestic interest rates and inflation but also by external stability indicators, such as foreign reserves and the exchange rate. The analysis offers valuable insights for policymakers and investors, suggesting that a comprehensive understanding of these relationships can enhance decision-making in both fiscal and monetary policy formulation, as well as investment strategies. As global economic conditions evolve, these macroeconomic relationships will continue to be crucial in navigating the challenges posed by inflation, financial market volatility, and external shocks.

References

- [1] Agrawal, N. (2020). Forecasting Recession in USA after Covid'19: Using Yield Curve Spread and Various Economic Factors, International Journal of Management, 11(11), 1455-1460.
- [2] Amin, M. R., & Mollick, A. V. (2022). Stock returns, oil prices and leverage: evidence from US firms. International Journal of Managerial Finance, 18(5), 785–811. https://doi.org/10.1108/IJMF-06-2021-0257
- [3] Bhar, R., & Malliaris, A. G. (2011). Dividends, Momentum, and Macroeconomic Variables as Determinants of the US Equity Premium Across Economic Regimes. Review of Behavioral Finance, 3(1), 27–53. https://doi.org/10.1002/rbf.14.
- [4] Çepni, O., Gül, S., Yılmaz, M. H., & Lucey, B. (2022). The impact of oil price shocks on Turkish sovereign yield curve. International Journal of Emerging Markets, 17(9), 2258–2277. https://doi.org/10.1108/IJOEM-06-2020-0681.
- [5] Chiang, T. C. (2022). Can gold or silver be used as a hedge against policy uncertainty and COVID-19 in the Chinese market? China Finance Review International, 12(4), 571–600. https://doi.org/10.1108/CFRI-12-2021-0232.
- [6] Dharani, M., Hassan, M. K., Huda, M., & Abedin, M. Z. (2023). Covid-19 pandemic and stock returns in India. Journal of Economics and Finance, 47(1), 251–266. https://doi.org/10.1007/s12197-022-09586-8.
- [7] Flannery, M. J., & Protopapadakis, A. A. (2002). Macroeconomic Factors Do Influence Aggregate Stock Returns. Review of Financial Studies, 15(3), 751–782. https://doi.org/10.1093/rfs/15.3.751.
- [8] Francová, B. (2018). An analysis of the impact of selected factors on the bond market. Acta Universitatis Agriculturae et Silviculturae Mendelianae Brunensis, 66(6), 1451–1458. https://doi.org/10.11118/actaun201866061451.
- [9] Garg, K., & Kalra, R. (2018). Impact of Macroeconomic Factors on Indian Stock Market. Parikalpana: KIIT Journal of Management, 14(1), 134. https://doi.org/10.23862/kiit-parikalpana/2018/v14/i1/173248.
- [10] Gupta, K., & Ahmed, S. (2020). Determinants of foreign portfolio flows to Indian debt market. Journal of Indian Business Research, 12(4), 459–479. https://doi.org/10.1108/JIBR-01-2019-0024.
- [11] Hughen, J. C., & Beyer, S. (2015). Stock returns and the US dollar: the importance of monetary policy. Managerial Finance, 41(10), 1046–1058. https://doi.org/10.1108/MF-09-2014-0234.
- [12] Hui, E. C. M., & Chan, K. K. (2022). How does Covid-19 affect global equity markets? Financial Innovation, 8(1). https://doi.org/10.1186/s40854-021-00330-5.
- [13] Jiang, R., Wang, B., Wu, C., & Zhang, Y. (2023). Impacts of Macroeconomic News Announcements on Corporate Bond Market (Vol. 11). https://doi.org/10.1108/s2514-465020230000011005.
- [14] Kalu, E., Okoyeuzu, C., Ukemenam, A., & Ujunwa, A. (2020). Spillover effects of the US monetary policy normalization on African stock markets. Journal of Economics and Development, 22(1), 3–19. https://doi.org/10.1108/jed-09-2019-0038.

- [15] Kariyawasam, N. P., & Jayasinghe, P. (2022). Determinants of sovereign spreads in Sri Lanka: global factors and countryspecific fundamentals. Asian Journal of Economics and Banking, 6(2), 236–254. https://doi.org/10.1108/ajeb-11-2021-0124
- [16] Koroleva, E., & Kopeykin, M. (2022). Understanding of Macro Factors That Affect Yield of Government Bonds. Risks, 10(8). https://doi.org/10.3390/risks10080166.
- [17] Kyereboah-Coleman, A., & Agyire-Tettey, K. F. (2008). Impact of macroeconomic indicators on stock market performance: The case of the Ghana Stock Exchange. Journal of Risk Finance, 9(4), 365–378. https://doi.org/10.1108/15265940810895025.
- [18] Laila, N., Rusmita, S. A., Cahyono, E. F., & Azman-Saini, W. N. W. (2021). The role of financial factors and non-financial factors on corporate bond and sukuk rating Indonesia. Journal of Islamic Accounting and Business Research, 12(8), 1077–1104. https://doi.org/10.1108/JIABR-10-2019-0187
- [19] Lakdawala, A., Pratap, B., & Sengupta, R. (2023). Impact of RBI's monetary policy announcements on government bond yields: evidence from the pandemic. In Indian Economic Review (Vol. 58, Issue s2). Springer India. https://doi.org/10.1007/s41775-023-00171-2.
- [20] Lebdaoui, H., Kiyadi, I., Bendriouch, F. Z., Chetioui, Y., Lebdaoui, F., & Alhayki, Z. (2024). The impact of COVID-19 stringency measures on emerging stock market stability: Does economic resilience matter? Journal of Economic and Administrative Sciences, 2019. https://doi.org/10.1108/JEAS-04-2023-0083.
- [21] Marisetty, N. (2024). Mitigating assumption violations in regression analysis: Insights from 20 years of currency pair data. International Journal of Financial Management and Economics, 7(2), 530–538. https://doi.org/10.33545/26179210.2024.v7.i2.412.
- [22] Michelson, N., & Stein, R. (2023). Factors Explaining Long-Term Government Bond Yields in the OECD Countries. SSRN Electronic Journal, March, 1–28. https://doi.org/10.2139/ssrn.4407123.
- [23] Moussa, F., & Delhoumi, E. (2022). The asymmetric impact of interest and exchange rate on the stock market index: evidence from MENA region. International Journal of Emerging Markets, 17(10), 2510–2528. https://doi.org/10.1108/IJOEM-01-2020-0089.
- [24] Olivares Rios, A., Rodríguez, G., & Ataurima Arellano, M. (2019). Estimation of Peru's sovereign yield curve: the role of macroeconomic and latent factors. Journal of Economic Studies, 46(3), 533–563. https://doi.org/10.1108/JES-04-2017-0089.
- [25] Ozcelebi, O., Perez-Montiel, J., & Manera, C. (2024). Bond yield spreads and exchange market pressure in emerging countries. International Journal of Emerging Markets. https://doi.org/10.1108/IJOEM-01-2023-0052.
- [26] Panigrahi, A., Karwa, P., & Joshi, P. (2019). Impact of macroeconomic variables on the performance of mutual funds: a selective study. Journal of Economic Policy & Research, 15(1), 29–43. https://ssrn.com/abstract=3653076.
- [27] Rao, A., Gupta, M., Sharma, G. D., Mahendru, M., & Agrawal, A. (2022). Revisiting the financial market interdependence during COVID-19 times: a study of green bonds, cryptocurrency, commodities, and other financial markets. International Journal of Managerial Finance, 18(4), 725–755. https://doi.org/10.1108/IJMF-04-2022-0165
- [28] Raza Rabbani, M., Hassan, M. K., Jamil, S. A., Sahabuddin, M., & Shaik, M. (2024). Revisiting the impact of geopolitical risk on Sukuk, stocks, oil and gold markets during the crises period: fresh evidence from wavelet-based approach. Managerial Finance, 50(3), 514–533. https://doi.org/10.1108/MF-12-2022-0587/
- [29] Sethy, T. K., & Tripathy, N. (2024). A new method for estimating liquidity and stock returns in Indian stock market. China Accounting and Finance Review, 26(2), 253–275. https://doi.org/10.1108/cafr-05-2023-0052.
- [30] Sreenu, N., & Pradhan, A. K. (2023). The effect of COVID-19 on Indian stock market volatility: can economic package control the uncertainty? Journal of Facilities Management, 21(5), 798–815. https://doi.org/10.1108/JFM-12-2021-0162/
- [31] Takahashi, T. (2022). Measuring Yield Curve Movements: A Principal Component Analysis of Spot Rate Changes in the JPY, USD, GBP, and CHF Interest Rate Swap Markets. 10, 193–208. https://doi.org/10.1108/s2514-465020220000010011.
- [32] Tangjitprom, N. (2012). The Review of Macroeconomic Factors and Stock Returns. International Business Research, 5(8). https://doi.org/10.5539/ibr.v5n8p107.
- [33] Trang Nguyen, H. Le, & Nguyen, P. A. (2022). The Impacts of Fiscal and Macroeconomic Factors on Vietnam Government Bond Yield. International Journal of Economics and Finance, 14(8), 23. https://doi.org/10.5539/ijef.v14n8p23.
- [34] Tripathi, V., & Seth, R. (2014). Stock Market Performance and Macroeconomic Factors: The Study of Indian Equity Market. Global Business Review, 15(2), 291–316. https://doi.org/10.1177/0972150914523599.
- [35] Verma, R. K., & Bansal, R. (2021). Impact of macroeconomic variables on the performance of stock exchange: a systematic review. International Journal of Emerging Markets, 16(7), 1291–1329. https://doi.org/10.1108/IJOEM-11-2019-0993/.
- [36] Yilanci, V., & Pata, U. K. (2023). COVID-19, stock prices, exchange rates and sovereign bonds: a wavelet-based analysis for Brazil and India. International Journal of Emerging Markets, 18(11), 4968–4986. https://doi.org/10.1108/IJOEM-09-2021-1465.
- [37] Zhou, S. (2021). Macroeconomic determinants of long-term sovereign bond yields in South Africa. Cogent Economics and Finance, 9(1). https://doi.org/10.1080/23322039.2021.1929678.
- [38] Zhou, Y., Teresiené, D., Keliuotyté-Staniuléniené, G., Kanapickiene, R., Dong, R. K., & Kaab Omeir, A. (2022). The Impact of COVID-19 Pandemic on Government Bond Yields. Frontiers in Environmental Science, 10(June), 1–18. https://doi.org/10.3389/fenvs.2022.881260.